

IEOR 6711: Stochastic Models I, Professor Whitt

SOLUTIONS to Homework Assignment 7

Problem 3.12 Take $h(t) = \mathbf{1}_{(0,a]}(t)$.

Problem 3.13 Since the state circulates the state space $\{1, 2, \dots, n\}$ in the same order. Hence we can define the alternating renewal process by *on* when it is in the state i and *off* when it is among $\{1, \dots, i-1, i+1, \dots, n\}$. Define $\mu_i = \int \bar{F}_i(t) dt$. Then

$$P(\text{process is in } i) \rightarrow \frac{E[\text{on}]}{E[\text{on}] + E[\text{off}]} = \frac{\mu_i}{\sum_{j=1}^n \mu_j}.$$

Problem 3.14 (a) $[t-x, t]$

(b) $[t, t+x]$

(c) $P(Y(t) > x) = P(A(t+x) > x)$

(d) See Problem 3.3.

Problem 3.15 (a)

$$\begin{aligned} P(Y(t) > x | A(t) = s) &= P(X_{N(t)+1} > x + s | \text{time at } t \text{ since the last renewal} = s) \\ &= \frac{\bar{F}(x+s)}{\bar{F}(s)}. \end{aligned}$$

(b) Using (a),

$$P(Y(t) > x | A(t+x/2) = s) = \begin{cases} 0 & \text{if } s < \frac{x}{2} \\ \frac{\bar{F}(s+x/2)}{\bar{F}(s)} & \text{if } s \geq \frac{x}{2} \end{cases}.$$

(c)

$$P(Y(t) > x | A(t+x) > s) = \begin{cases} 1 & \text{if } s \geq x \\ P(\text{no events in } [t, t+x-s]) = e^{-\lambda(x-s)} & \text{if } s < x \end{cases}.$$

(d)

$$P(Y(t) > x, A(t) > y) = P(Y(t-y) > x+y) = P(A(t+x) > x+y).$$

(e)

$$\frac{A(t)}{t} = \frac{t - S_{N(t)}}{t} = 1 - \frac{S_{N(t)}}{N(t)} \frac{N(t)}{t} \rightarrow 1 - \mu \frac{1}{\mu} = 0.$$

Problem 3.16

$$\begin{aligned}
\lim_{t \rightarrow \infty} \mathbb{E}[Y(t)] &= \frac{\mathbb{E}[X^2]}{2\mu} \\
&= \frac{n \left(\frac{1}{\lambda}\right)^2 + \left(\frac{n}{\lambda}\right)^2}{2\frac{n}{\lambda}} \\
&= \frac{1+n}{2\lambda}.
\end{aligned}$$

To get it without any computations, consider a Poisson process with rate λ and say the a *renewal* occurs at the Poisson events numbered $n, 2n, \dots$. Now at time t , t large, it is equally likely that the most recent event was an event of the form $i+kn$, $i = 0, 1, 2, \dots, n-1$. That is, modulo n , the number of the most recent Poisson event is equally likely to be $n, 1, \dots, n-1$. Conditioning on the value of this quantity gives that for the renewal process

$$\lim_{t \rightarrow \infty} \mathbb{E}[Y(t)] = \frac{1}{n} \left(\frac{1}{\lambda} + \dots + \frac{n}{\lambda} \right) = \frac{n+1}{2\lambda}.$$

Problem 3.18 (a) Delayed renewal process.

(b) Neither.

If F is exponential,

(a) Delayed renewal process.

(b) Renewal process.

Problem 3.21 Let X_i equal 1 if the gambler wins bet i , and let it be 0 otherwise. Also, let N denote the first time the gambler has won k consecutive bets. Then $X = \sum_{i=1}^N X_i$ is equal to the number of bets that he wins, and $X - (N - X) = 2X - N$ is his winnings. By Wald's equation

$$\mathbb{E}[X] = p\mathbb{E}[N] = p \sum_{i=1}^k p^{-i}.$$

Thus

$$(a) \mathbb{E}[2X - N] = 2\mathbb{E}[X] - \mathbb{E}[N] = (2p - 1)\mathbb{E}[N] = (2p - 1) \sum_{i=1}^k p^{-i}$$

$$(b) \mathbb{E}[X] = p \sum_{i=1}^k p^{-i}$$

Problem 3.22 (a)

$$\begin{aligned}
\mathbb{E}[T_{HHHTTHH}] &= \mathbb{E}[T_{HH}] + p^{-4}(1-p)^{-2} \\
&= \mathbb{E}[T_H] + p^{-2} + p^{-4}(1-p)^{-2} \\
&= p^{-1} + p^{-2} + p^{-4}(1-p)^{-2}
\end{aligned}$$

(b) $E[T_{HTHTT}] = p^{-2}(1-p)^{-3}$

$E[N_{B|A}] = E[N_{HTHTT|H}] = E[N_{HTHTT}] - E[N_H] = 32 - 2 = 30$, $E[N_{A|B}] = E[N_A] = 64 + 4 + 2 = 70$ and $E[N_B] = 32$.

(c) $P_A = (32 + 70 - 70)/(30 + 70) = 0.32$

(d) $E[M] = 32 - 30(0.32) = 22.4$

Problem 3.23 Let H denote the first k flips and Ω is the set of all possible H . Conditioning on H gives:

$$\begin{aligned} E[\text{number until repeat}] &= \sum_{H \in \Omega} E[\text{number until repeat}|H]P(H) \\ &= \sum_{H \in \Omega} \frac{1}{P(H)}P(H) = |\Omega| = 2^k \end{aligned}$$

Problem 3.25 (a) First note that

$$E[N_D(t)|X_1 = x] = \begin{cases} 1 + E[N(t-x)] & \text{if } x \leq t \\ 0 & \text{if } x > t \end{cases}$$

$$\begin{aligned} m_D(t) = E[N_D(t)] &= \int_0^\infty E[N_D(t)|X_1 = x]dG(x) \\ &= \int_0^t (1 + E[N(t-x)])dG(x) \\ &= G(t) + \int_0^t m(t-x)dG(x) \end{aligned}$$

(b)

$$\begin{aligned} E[A_D(t)] &= E[A_D(t)|S_{N_D(t)} = 0]\bar{G}(t) + \int_0^t E[A_D(t)|S_{N_D(t)} = s]\bar{F}(t-s)dm_D(s) \\ &= t\bar{G}(t) + \int_0^t (t-s)\bar{F}(t-s)dm_D(s) \\ &\xrightarrow{t \rightarrow \infty} \frac{1}{\mu} \int_0^\infty t\bar{F}(t)dt \quad \text{By key renewal theorem (Proposition 3.5.1(v))} \\ &= \frac{1}{\mu} \int_0^\infty t \int_t^\infty dF(s)dt \\ &= \frac{\int_0^\infty s^2 dF(s)}{2 \int_0^\infty s dF(s)} \end{aligned}$$

(c) $t\bar{G}(t) = t \int_t^\infty dG(x) \leq \int_t^\infty s dG(s) \xrightarrow{t \rightarrow \infty} 0$ since $\int_0^\infty s dG(s) < \infty$.

(Here we used the so-called *dominated convergence theorem*.)

$$\begin{aligned} \int_n^\infty s dG(s) &= \int_0^\infty s \mathbf{1}_{[n, \infty)}(s) dG(s) \\ &\xrightarrow{n \rightarrow \infty} \int_0^\infty \lim_{n \rightarrow \infty} s \mathbf{1}_{[n, \infty)}(s) dG(s) = \int_0^\infty 0 dG(s) \end{aligned}$$

dominated convergence theorem

since $s\mathbf{1}_{[n,\infty)}(s) \leq s$ and s is integrable with respect to $G(\cdot)$ from $\int_0^\infty s dG(s) < \infty$ and $s\mathbf{1}_{[n,\infty)}(s) \rightarrow 0$ for each s in *pointwise* sense. (Check the conditions for the dominated convergence theorem.) Now we extend n to t using monotonicity of the integral. Wow! This is a good example showing that if you are familiar with a little rigorous *analysis*, then it's O.K. with only one line. But if not, you should practice the underlying logic whenever you encounter them.)

Problem 3.28 Using the *uniformity* of each Poisson arrival under given $N(t)$,

$$E[\text{Cost of a cycle}|N(T)] = K + N(T) \times c \times \frac{T}{2}$$

and so

$$\frac{E[\text{Cost}]}{E[\text{Time}]} = \frac{K + \lambda c T^2 / 2}{T} = \frac{K}{T} + \frac{\lambda c T}{2}$$

which is minimized at $T^* = \sqrt{2K/\lambda c}$ and minimal average cost is thus $\sqrt{2\lambda K c}$. On the other hand the optimal value of N is (using calculus) $N^* = \sqrt{2\lambda K/c}$ and the minimal average cost is $\sqrt{2\lambda c K} - \frac{c}{2}$.