

Lecture VIII

Ito's lemma: stochastic calculus

Ito's lemma

• Heath - Jarrow - Merton model.

↳ Martingale

↳ Arbitrage free models

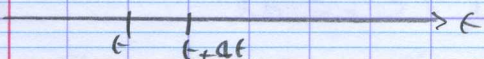
↳ relationship between Martingale & Arbitrage-free

$S(t)$ = price of asset at time t .

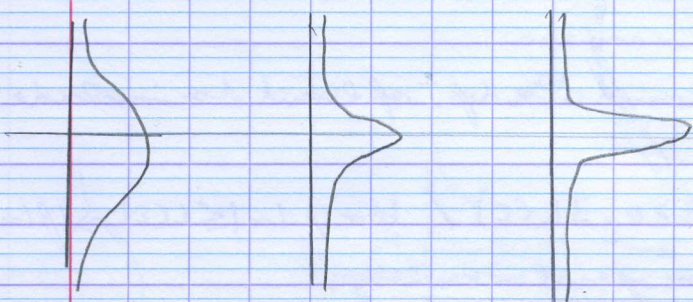
ex: price of Merrill Lynch stock.

$G_t(S(t), t)$ = price of contract based on t and $S(t)$

$dW(t) = N(0, \Delta t)$, Δt = variance.



$$\Delta S = S(t+\Delta t) - S(t) = \mu(t, S) \Delta t + \sigma(t, S) dW(t)$$



$\begin{cases} \mu(t, S) = \text{drift} \\ \sigma(t, S) = \text{volatility} \end{cases}$

$$dS = \mu(t, S) dt + \sigma(t, S) dW(t)$$

For ease we use, $\begin{cases} \mu(t, S) = \mu \\ \sigma(t, S) = \sigma \end{cases}$

What is $dG_t = G_t(S(t+\Delta t), t+\Delta t) - G_t(S(t), t)$?

What is dG_t ?

Recall Taylor's expansion:

$$f(x+\Delta x) = f(x) + \Delta x f'(x) + \frac{\Delta x^2}{2} f''(x) + \dots$$

$$\Delta f = f(x+\Delta x) - f(x) = \Delta x f'(x) + \frac{\Delta x^2}{2} f''(x)$$

$$\Delta f \xrightarrow{\Delta t \rightarrow 0} df = f'(x)$$

ex:

$$\begin{cases} f(x) = x^2 \\ f(1) = 1 \\ f(1.01) = 1.2001 \end{cases}$$

$$f(1.01) - f(1) = 0.2$$

$$\boxed{df = (2 \cdot x) dx}$$

$$G(x, y) = \dots$$

$$G(x+\Delta x, y+\Delta y) = G(x, y) + (\Delta x \frac{\partial}{\partial x} + \Delta y \frac{\partial}{\partial y}) G + \frac{1}{2!} (\Delta x^2 \frac{\partial^2}{\partial x^2} + 2\Delta x \Delta y \frac{\partial^2}{\partial x \partial y} + \Delta y^2 \frac{\partial^2}{\partial y^2}) G$$

$$\Delta G = G(x+\Delta x, y+\Delta y) - G(x, y)$$

$$= \Delta x \frac{\partial G}{\partial x} + \Delta y \frac{\partial G}{\partial y} + \frac{1}{2} \Delta x^2 \frac{\partial^2 G}{\partial x^2} + \Delta x \Delta y \frac{\partial^2 G}{\partial x \partial y} + \frac{\Delta y^2}{2} \frac{\partial^2 G}{\partial y^2}$$

$$\Delta x \rightarrow 0, \Delta y \rightarrow 0$$

$$dG = \frac{\partial G}{\partial t} dt + \frac{\partial G}{\partial S} dS + \frac{1}{2} \left[\frac{\partial^2 G}{\partial F^2} dt^2 + 2 dS dt \frac{\partial^2 G}{\partial t \partial S} + \frac{\partial^2 G}{\partial S^2} dS^2 \right]$$

Try avoid every term of an higher power than dt.

$$\rightarrow dG = \frac{\partial G}{\partial t} dt + \frac{\partial G}{\partial S} (\mu dt + \sigma dW_t) + \frac{1}{2} \frac{\partial^2 G}{\partial S^2} dS^2$$

because $dt^2 \ll dt$,

$$dS \sim \sqrt{dt} \Rightarrow |dS dt| \ll dt^{3/2} \ll dt$$

$$dS \approx dt$$

$$dS^2 = (\mu dt + \sigma dW_t)^2$$

$$= \mu^2 dt^2 + 2\mu\sigma dt dW_t + \sigma^2 dW_t^2$$

$$\approx \sigma^2 dW_t^2$$

def: We have to extend the defⁿ of limit to include random variables.

$$\lim_{x \rightarrow 0} f(x) = A \Leftrightarrow \forall \epsilon > 0, \exists \delta(\epsilon) / \forall x, |x| \leq \delta(\epsilon) \Rightarrow |f(x) - A| \leq \epsilon$$

Convergence in probability:

$\{X_n\}$: random variables,

then: $X_n \rightarrow X$ in probability

$$\Leftrightarrow \forall \epsilon > 0, \lim_{n \rightarrow \infty} P[\omega, |X_n(\omega) - X(\omega)| \geq \epsilon] = 0$$

Let's prove that $(dW_t)^2 \approx dt$.

$$W_t \sim N(0, t)$$

$$dW_t \sim N(0, dt)$$

$$dW_t \sim \sqrt{dt} N(0, 1)$$

$$(dW_t)^2 \sim dt (N(0, 1))^2$$

$$E[(dW_t)^2] = dt E[(N(0, 1))^2]$$

$$= dt \text{Var}[N(0, 1)] = dt$$

$$\begin{aligned}
 \text{Var}[(dW_t)^2] &= \text{Var}[dt (N(0,1))^2] \\
 &= dt^2 \text{Var}[N(0,1)^2] \\
 &= dt^2 [E[N(0,1)^2] - (E[N(0,1)^2])^2] \\
 &= dt^2 (3 - 1) \\
 &= 2 dt^2
 \end{aligned}$$

$$\begin{aligned}
 \text{Var}[(dW_t)^2] &\xrightarrow{dt \rightarrow 0} 0 \\
 \Rightarrow (dW_t)^2 &\xrightarrow{dt \rightarrow 0} E[(dW_t)^2]
 \end{aligned}$$

$$\Rightarrow \boxed{(dW_t)^2 \xrightarrow{dt \rightarrow 0} dt}$$

\Rightarrow Itô's Lemma:

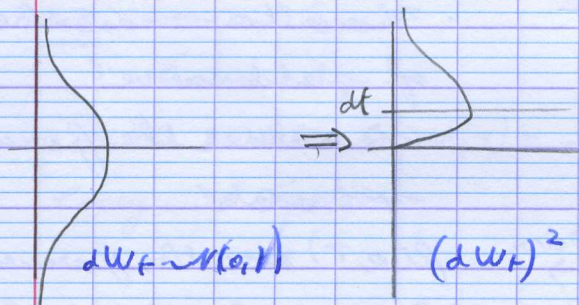
$$dG(S(t), t) = \left[\frac{\partial G}{\partial t} + \mu \frac{\partial G}{\partial S} + \frac{\sigma^2}{2} \frac{\partial^2 G}{\partial S^2} \right] dt + \sigma \frac{\partial G}{\partial S} dW_t$$

Recall if $X \sim N(0,1)$ then: $E[X] = 0$

$$E[X^2] = \text{Var}(X) = 1$$

$$\text{skewness: } E[X^3] = 0$$

$$\text{kurtosis: } E[X^4] = 3$$



example:

$$\textcircled{1} G(S(t), t) = S(t)^2$$

$$\text{with } dS(t) = \mu dt + \sigma dW_t$$

$$G(x, t) = G(x) = x^2$$

$$\text{using: } dG = \left[\frac{\partial G}{\partial t} + \mu S \frac{\partial G}{\partial S} + \frac{\sigma^2 S^2}{2} \frac{\partial^2 G}{\partial S^2} \right] dt + \sigma \frac{\partial G}{\partial S} S dW_t$$

$$\frac{\partial G}{\partial t} = 0$$

$$\frac{\partial G}{\partial S} = 2x$$

$$\frac{\partial^2 G}{\partial S^2} = 2$$

$$Y(t) = S(t)^2$$

$$dY(t) = \left[\mu \times 2S(t) + \frac{\sigma^2}{2} \times 2S(t)^2 \right] dt + \sigma \times 2S(t) dW_t$$

$$= (2\mu S(t) + \sigma^2 S(t)^2) dt + 2\sigma Y(t) dW_t$$

def: Chebyshev's Inequality

X is a random variable with mean μ , variance σ^2

$$P[|X - \mu| \geq \epsilon] \leq \frac{\text{Var}(X)}{\epsilon^2}$$

rem: if we prove that $E(X) \rightarrow \mu$
and $\text{Var}(X) \xrightarrow{dt \rightarrow 0} 0$

we have $P[|X - \mu| \geq \epsilon] \xrightarrow{dt \rightarrow 0} 0$

and we obtain the convergence in probability thanks to Bienaymé - Chebyshev's Inequality!

• for $x = dW(t)^2$

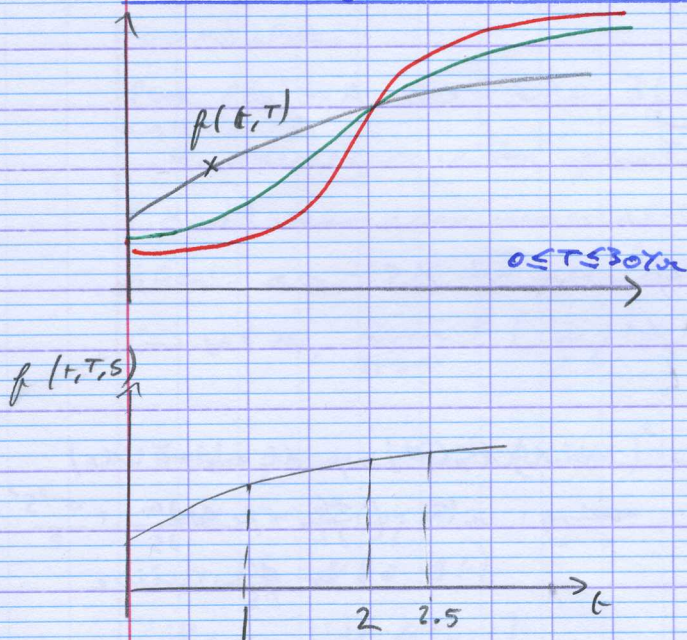
$$E[x^2] = E[dW(t)^4] = 3dt^2$$

$$E[x] = E[dW(t)^2] = dt$$

$$\text{Var}[x] = \text{Var}[dW(t)^2] = 3dt^2 - dt^2 = 2dt^2 \xrightarrow{dt \rightarrow 0} 0$$

$$\Rightarrow P[|dW(t)^2 - dt| \geq \epsilon] \xrightarrow{dt \rightarrow 0} 0$$

Heath - Jarrow - Morton Model.



• How to model the movements of yield curve?

• We model the forward spot curve.

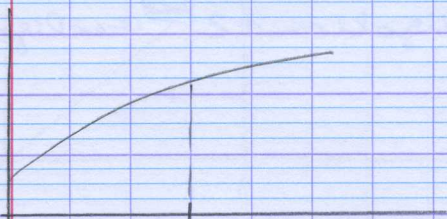
• $f(0, T)$ are the forward rates observed today.

or equivalently we can model the movement of 0-coupon bond prices $P(t, T)$

recall $f(t, T) = - \frac{\partial \log P(t, T)}{\partial T}$

$$df(t, T) = \mu(t, T) dt + \sum_{i=1}^n \sigma_i(t, T, f(t, T)) dW_i(t)$$

and $dW_i(t)$ are independent Wiener processes.



Can $\mu(t, T)$ and $\sigma_i(t, T, \mu(t, T))$ take any form?

Arbitrage:

consider $t=0$ and suppose we have n assets A_1, \dots, A_n , with $V_1(0), \dots, V_n(0)$.

Let $\sum_{i=1}^n \alpha_i V_i(0)$ be a portfolio of these assets.

Let $V(t) = \sum_{i=1}^n \alpha_i V_i(t)$ be the value of this portfolio at time t which is random.

There is an arbitrage $\Leftrightarrow V(t) \geq 0$ for all realization of prices at time t and $V(t) > 0$ for some realisation

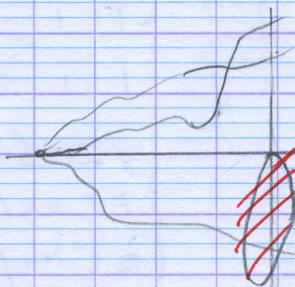
ie: $\begin{cases} P(V(t) \geq 0) = 1 \\ P(V(t) > 0) > 0 \end{cases}$

ex: GM 36.

Kard 10

sell 10 GM $\cdot V(0) = -10 \text{ GM} + 36 \text{ €} = 0$

buy 36 Kard



def: X_1, X_2 is a martingale $\Leftrightarrow E[X_{i+1} | X_i] = X_i$

pty: $E[X_{i+1}] = E[X_i]$

proof:

$$E[X_{i+1} | X_i] = X_i \Rightarrow E[E[X_{i+1} | X_i]] = E[X_i]$$

$$\Rightarrow E[X_{i+1}] = E[X_i]$$

$V(0), V(1)$

$$E[V(1)] = \sum \text{prob. (prices)} \times \text{prices}$$

Suppose V is a martingale $E[V(1) | V(0)] = V(0) = 0$

\forall values of $V(1) = 0$.

| | | | |
|-------|-------|-----|---|
| S_1 | S_2 | | $X_1 = 1^{st}$ realization = S_1 |
| 0 | 0 | 0.5 | $X_2 = 1^{st}$ realization + 2^{nd} realization |
| 1 | 0 | 0.3 | = $S_1 + S_2$. |
| 0 | 1 | 0.1 | |
| 1 | 1 | 0.1 | |

• If $X_1 = 0$, $E[X_2 | X_1] = E[X_2 | X_1 = 0]$
 $= 0 \cdot \frac{0.5}{0.6} + 1 \cdot \frac{.1}{.6} = \frac{1}{6}$

• $E[X_2 | X_1 = 1] = 1 \cdot \frac{.3}{.4} + 2 \cdot \frac{.1}{.4}$

Under this probability the process is not a martingale.

Is it possible to change the probabilities (measure) such that, under this new probability the process is a martingale.

Are these probabilities equivalent?

| | old measure | New prob. |
|-------|-------------|-----------|
| -1 -1 | .5 | .25 |
| 1 -1 | .3 | .25 |
| -1 1 | .1 | .25 |
| 1 1 | .1 | .25 |

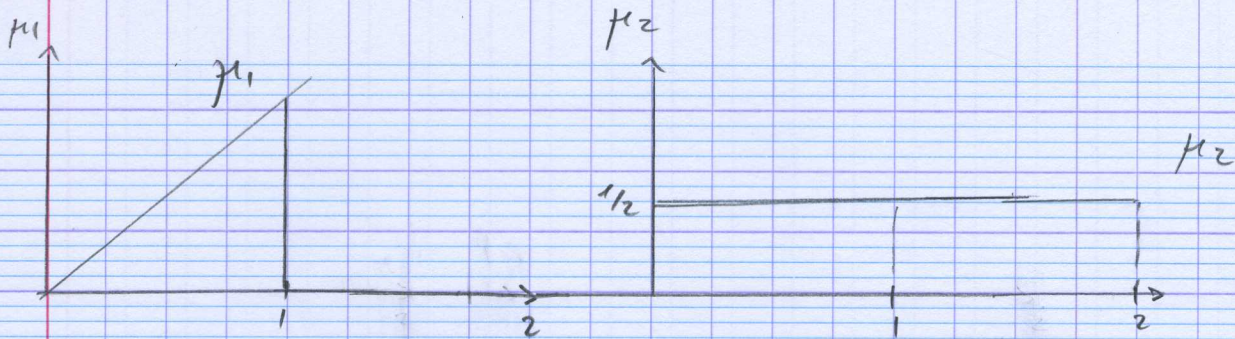
proof:

$$\begin{cases} E[X_2 | X_1 = -1] = -2 \times \frac{1}{2} + 0 \times \frac{1}{2} = -1 \\ E[X_2 | X_1 = 1] = 0 \times \frac{1}{2} + 2 \times \frac{1}{2} = 1 \end{cases}$$

• Probability space $(\mathcal{F}, \mathcal{F}, \mu)$

$$\begin{pmatrix} -1 & -1 \\ 1 & -1 \\ -1 & 1 \\ 1 & 1 \end{pmatrix}$$

def: μ_1 and μ_2 are equivalent $\Leftrightarrow A \rightarrow \mu_1(A) = 0 \Leftrightarrow \mu_2(A) = 0$.



$$f(x) = \begin{cases} 2x, & 0 \leq x \leq 1 \\ 0, & 1 \leq x \leq 2 \end{cases}$$

$$f(x) = \begin{cases} \frac{1}{2}, & 0 \leq x \leq 2 \\ 0, & x > 2 \end{cases}$$

$$\mu_1 p_1 [1.5, 2] = 0.$$

→ μ_1 and μ_2 are not equivalent.

Radon-Nikodym Theorem.

If μ_1 and μ_2 are equivalent
then: there are functions that enable us to calculate the probability of one event using μ_1 that is the same as the other over μ_2 .