

IEOR 4106 lec 16

CTMCs, PASTA

" $l = \lambda w$ " (Little's Law)

Global Balance

Equations for

CTMCs

rate out of state j

$$a_j p_j =$$

$$\sum_{i \neq j} a_i p_i p_{ij}$$

rate into state j

$\{p_j\}$ limiting probabilities

$$H_i \sim \exp(a_i) \quad i \in \mathcal{S}$$

$$p_j \stackrel{\text{def}}{=} \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t \mathbb{I}\{X(s) = j\} ds$$

$$j \in \mathcal{S}$$

$$T_{jj} = \begin{cases} \text{return time to state } j \\ | X(0) = j \\ \stackrel{\text{def}}{=} \infty \quad \text{if no return} \end{cases}$$

$$\geq H_j \quad \left(T_{jj} = \min\{n \geq 1 : X_n = j \mid X_0 = j\} \right)$$

$$\stackrel{\text{def}}{=} \infty \quad \text{if no return}$$

(we are interested in the case
 when $P(T_{jj} < \infty) = 1$
recurrence of state j)

$$1 = P(T_{jj} < \infty) \iff P(\tau_{jj} < \infty) = 1$$

j is recurrent

transient otherwise

II

Pos. rec. in
Cont. time
($X(t)$)

$$E(T_{jj}) < \infty$$

If j is recurrent,
then by the Markov property

visits to state j form a

renewal process $\{t_n(j)\}$

$t_n(j)$ = time at which $(X(t))$
visits state j for n^{th} time
 $n \geq 1$

(i) interarrival times

$\sim T_{jj}$

$(N_j^+(t))$ counting
process

$\{Y_m : m \geq 1\}$
 $Y_m = t_m(j) - t_{m-1}(j)$

$N_j^-(t) =$ number of times $(X(t))$
departs state j
during $(0, t)$

$N_j^+(t) =$ number of times $(X(t))$
enters state j
during $(0, t)$

they differ by at most **1**
for each t

$$\lim_{t \rightarrow \infty} \frac{N^+(t)}{t} = \frac{1}{E(T_{ij})} \quad \text{wpl}$$

$$= \frac{E R T}{E R T}$$

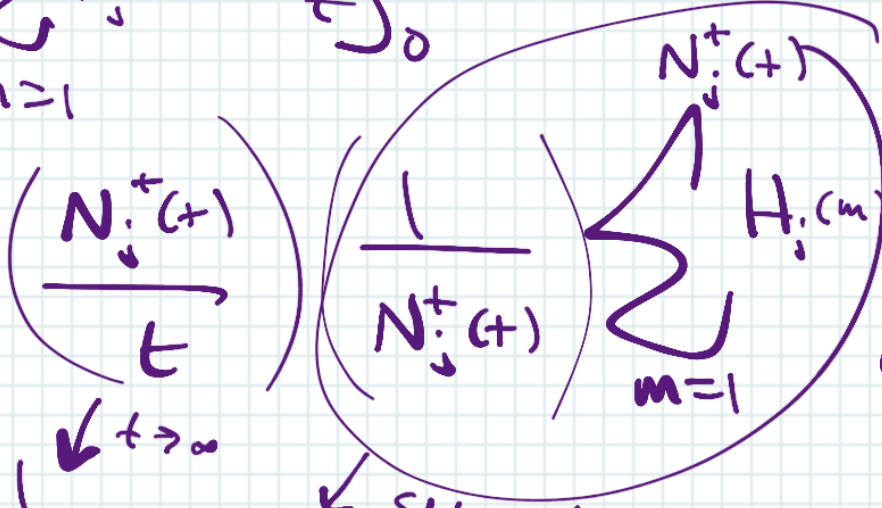
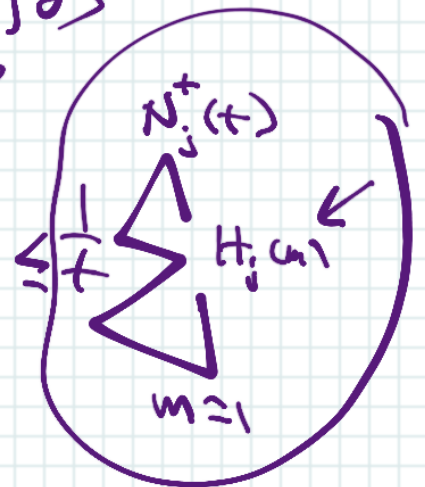
$$= \lim_{t \rightarrow \infty} \frac{N^-(t)}{t} \quad \text{wpl}$$

"rate $(X(t))$ enters j = rate $(X(t))$ leaves state j "

$$= \frac{1}{E(T_{ij})}$$

$$P_j = \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t \mathbb{I}\{X(s) = j\} ds$$

$$\frac{1}{t} \sum_{m=1}^{N_j^-(t)} H_j(m) \leq \frac{1}{t} \int_0^t \mathbb{I}\{X(s) = j\} ds$$



$$H_j(m) \stackrel{\text{dist}}{=} H_j$$

iid holding times per visit to j

ERT $\frac{1}{E(T_{jj})}$ $E(H_j)$ SLLN $t \rightarrow \infty$

expl a_j
 $E(H_j) = \frac{1}{a_j}$

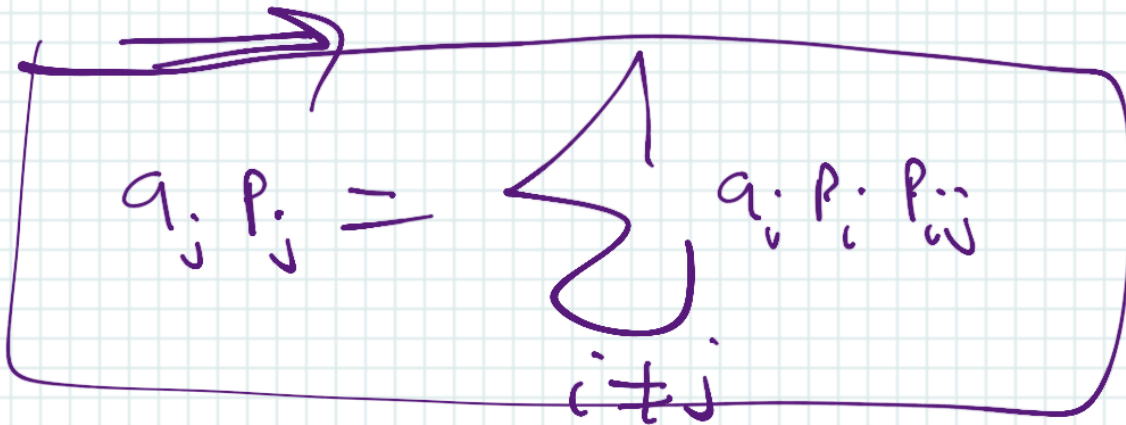
$$= \frac{E(H_j)}{E(T_{jj})} = \frac{1}{a_j}$$

left-hand side is the same

$$\Rightarrow p_j = \frac{1}{a_j} \quad j \in J$$

$$\Rightarrow \left(a_j p_j = \frac{1}{E(T_{jj})} = \text{rate out/in to state } j \right)$$

$$P_{01} = 1$$


$$a_j P_j = \sum_{i \neq j} a_i P_i P_{ij}$$

$$j \in \mathcal{S}$$

B & D process

$$a_j = \lambda_j + \nu_j \quad j \geq 0$$

$$(P_{i,i+1} + P_{i,i-1} = 1)$$

$$P_{i,i+1} = \frac{\lambda_i}{\lambda_i + \nu_i}$$

$$P_{i,i-1} = \frac{\nu_i}{\lambda_i + \nu_i}$$

$$\begin{aligned} \lambda_0 P_0 &= (\lambda_1 + \nu_1) P_1 \cdot P_{10} \\ &= \lambda_1 + \nu_1 P_1 \frac{\nu_1}{\lambda_1 + \nu_1} = \nu_1 P_1 \end{aligned}$$

$$\lambda_0 P_0 = \nu_1 P_1$$

$$(\lambda_1 + \nu_1) P_1 = \lambda_0 P_0 + (\nu_2 + \lambda_2) P_2 \quad \frac{\nu_2}{\lambda_2 + \nu_2} P_{2,1}$$

$$(\lambda_1 + \nu_1) P_1 = \lambda_0 P_0 + \nu_2 P_2$$

$$\Rightarrow \lambda_1 P_1 = \nu_2 P_2$$

$$(\lambda_2 + \nu_2) P_2 = \lambda_1 P_1 + \nu_3 P_3$$

$$\Rightarrow \lambda_2 P_2 = \nu_3 P_3 \dots$$

$$\lambda_j P_j = \nu_{j+1} P_{j+1} \quad j \geq 0$$

Birth & Death Balance Equation

// rate from $i \rightarrow i+1$ = rate from $i+1 \rightarrow i$

$j \leftarrow A$

$$\lambda_j p_j = \mu_{j+1} p_{j+1}, \quad j \geq 0$$

For a queueing model

$X(t)$ = number of customers in the system at time t

arrivals form a point process

$\{t_n\} \leftarrow$

$$P_j = \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t I\{X(s) = j\} ds$$

$$\pi_j^a = \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N I\{X(t_n^-) = j\}$$

= long-run proportion of customers j who arrive finding already in system

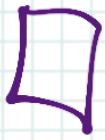
PASTA (Poisson Arrivals See
Time Averages)

if $\{t_n\}$ is a $PP(\lambda)$ (Poisson)

then $\underline{P_j = \pi_j^a}$, $j \in \mathcal{J}$

If arrivals are not Poisson, result
does not hold in general

Counterexample: $(X(t))$

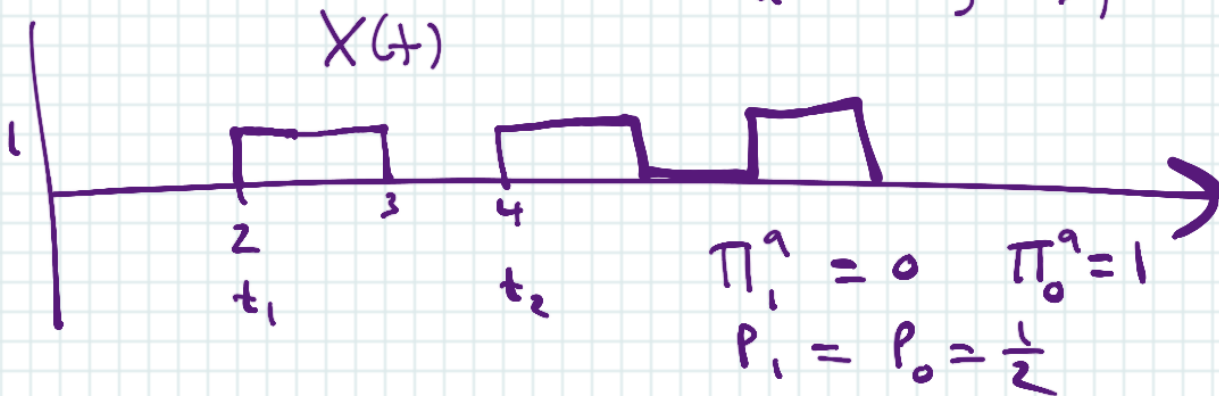


Single-Server Queue

$$S_n \equiv 1, n \geq 1$$

interarrival times $\equiv 2$

$$t_n = 2n, n \geq 1$$



Proof of PASTA for M/M/1 queue

rate at which
 $X(t)$ moves $j \rightarrow j+1$

$$\cancel{\lambda} \pi_j^a = \cancel{\lambda} p_j$$

basic
principles

Birth and
Death
Balance
Equations

$$\pi_j^a = p_j \quad j \geq 0$$

Nice application:

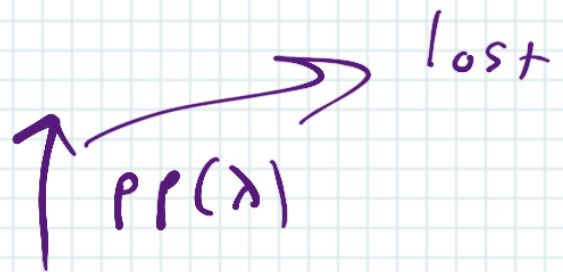
Erlang's Loss formula

M/M/c loss $\sum_{n=c}^{\infty} p_n$

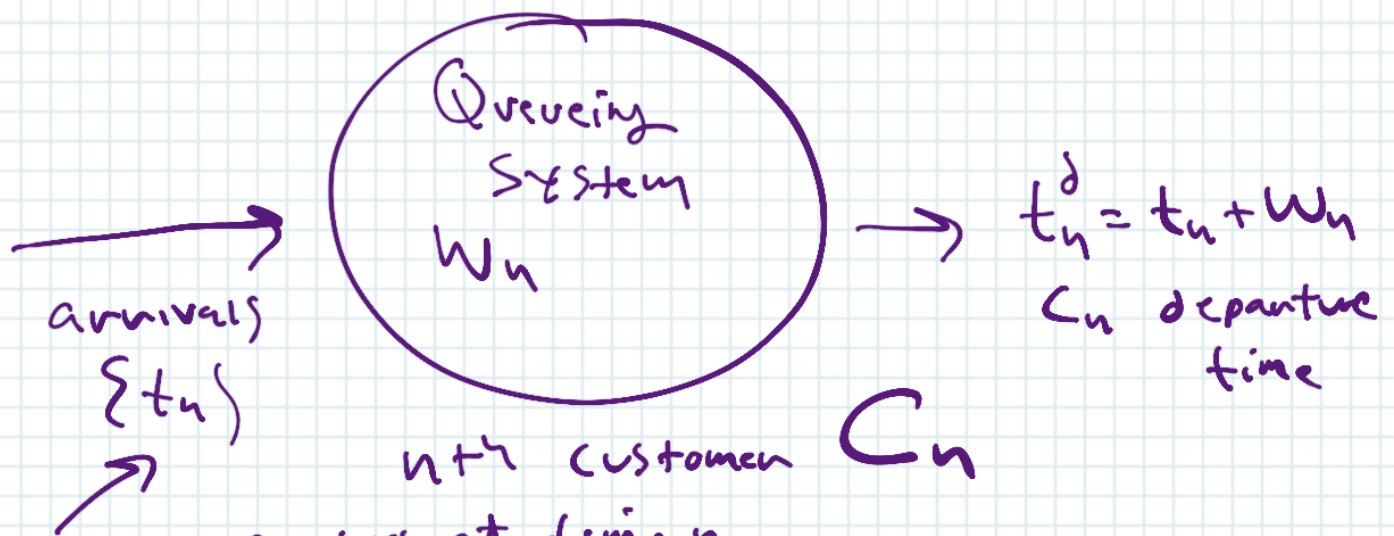
$$\pi_c^q \stackrel{\text{PASTA}}{=} p_c$$



Solved using
Balance equations



" $l = \lambda w$ " Little's Law



arrives at time n
and enters the system

has a sojourn time $= W_n \geq 0$

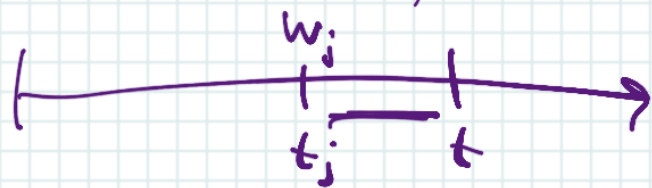
$=$ total time spent in the system

Defined if they exist

$$\lambda \stackrel{\text{def}}{=} \lim_{t \rightarrow \infty} \frac{N(t)}{t} \quad \underline{\text{arrival rate}}$$

$$W \stackrel{\text{def}}{=} \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n W_i \quad \text{average service time}$$

$L(t) =$ number of customers in the system at time t

$$= \sum_{i=1}^{N(t)} I\{W_i > t - t_j\}$$


$$l \stackrel{\text{def}}{=} \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t L(s) ds$$

time average
number of customers
in the system

" $l = \lambda w$ ": If λ and w exist
and are $< \infty$, then
 l exists and

$$l = \lambda w.$$

Example: SEAS
undergraduates

$$\lambda = 1500 \text{ year}$$

$$w = 4 \text{ years}$$

$$l = \lambda w = 6000 \text{ students}$$

Proof:

$$\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t L(s) ds$$

area
under
 $L(s)$
in $[0, t]$

$$\int_0^t L(s) ds$$

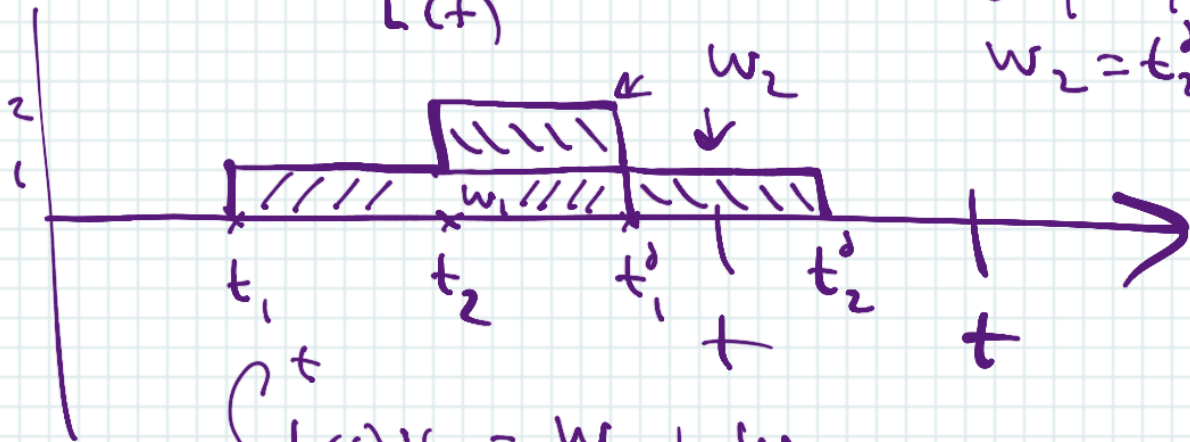
\sim
 \nearrow

$$\sum_{j=1}^{N(t)} w_j$$

$L(t)$

w_2

$$w_1 = t_1' - t_1$$
$$w_2 = t_2' - t_2$$



$$\int_0^t L(s) ds = w_1 + w_2$$

$$\frac{1}{t} \sum_{j: t_j^d \leq t} w_j$$

$$\leq \frac{1}{t} \int_0^t L(s) ds$$

$$\leq \frac{1}{t} \sum_{j=1}^{N(t)} w_j$$

$$\left(\frac{N(t)}{t} \right) \left(\frac{1}{N(t)} \sum_{j=1}^{N(t)} w_j \right)$$

λ
 w

λ, w exist and $< \infty$ by assumption

Left-hand side also
converses to $\lambda \omega$ (more complicated
proof)

So the result is complete.

$$l = \lambda \omega$$

Applications

∞ Server Queue

$(M/G/\infty \text{ queue})$



\uparrow $\{t_n\}$ λ $W_n = S_n$

$$\frac{1}{N} \stackrel{\text{def}}{=} w = \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{j=1}^N S_j$$

$S \stackrel{\text{def}}{=} \frac{\lambda}{N}$ average service time

if $\lambda, \frac{1}{\lambda}$ exist and $< \infty$,
then l exists

$$l = \rho$$

Recall $M|b| \infty$? never

$$l = \rho$$

$$X(t) \sim \text{Poisson}(d(t))$$

$$\lim_{t \rightarrow \infty} d(t) = \rho$$

$$E(X(t)) \rightarrow \rho$$

$$X(t) \Rightarrow \text{Poisson}(\rho)$$

