

IEOR 4106 Lec 17

- 1) More applications of " $l = \lambda w$ " (Little's Law)
- 2) forwards/backwards rec. time for renewal processes.
the inspection paradox

$$\underline{l = \lambda w}$$

$$l \stackrel{\text{def}}{=} \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t L(s) ds = \text{average number of customers in the system}$$

$$w \stackrel{\text{def}}{=} \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{j=1}^n w_j = \text{average sojourn time}$$

$$\lambda \stackrel{\text{def}}{=} \lim_{t \rightarrow \infty} \frac{N(t)}{t} = \text{arrival rate of customers entering the system}$$

If both λ and w exist and are $< \infty$, then l exists and $l = \lambda w$

examples:

1) ∞ -server queue $(t_n) (N(t))$

arrival rate λ

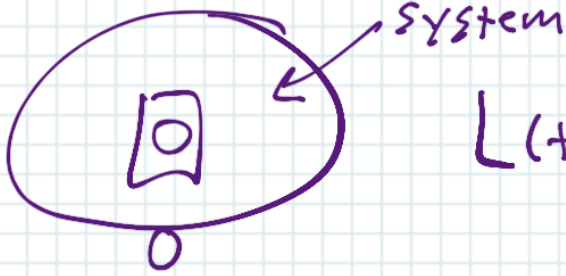
$W_i = S_i$, service times

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n W_i$$
$$= \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n S_i = \frac{1}{\mu}$$

$$S = \frac{\lambda}{\mu}$$

$$l = \lambda w = \lambda \cdot \frac{1}{\mu} = S$$

↳ Single-Server Queue (Let's use the Server as the system)



$$L(t) = \begin{cases} 1 & \text{if server is busy} \\ 0 & \text{if server is idle} \end{cases}$$

l = long-run proportion of time the server is busy

$\uparrow \{t_n\}$ λ = arrival rate

$\left(\int_0^t L(s) ds = \right.$
amount of time during $(0, t)$ that the server is busy/

$$W_i = S_i, \quad W = \frac{1}{\lambda}$$

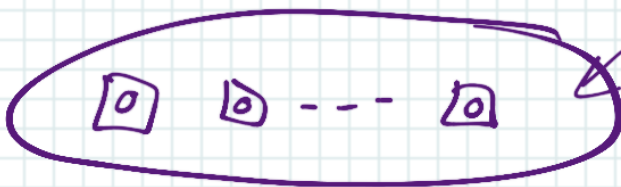
$$l = \lambda \cdot \frac{1}{\mu} = \rho$$

$$\left(\begin{array}{l} \rho < 1 \\ \lambda < N \end{array} \right)$$

$1 - \rho =$ long-run prop time the server is idle

3)

C-server queue (all C-servers are the system)



$L(t) =$ number of busy servers at t

$$\in \{0, 1, \dots, c\}$$

$$w_j = S_j$$

$$w = \frac{1}{\mu}$$

0
0
0 {t_n} rate λ
↑

$$\lambda < cN$$

$$(\rho < c)$$

$$l = \lambda \cdot \frac{1}{\mu} = \rho$$

ρ = long-run average number
of busy servers

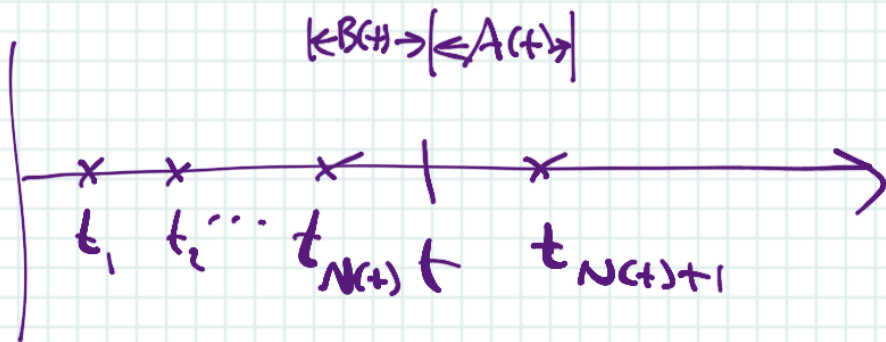
$$0 < \rho < 1$$

Renewal process $\{t_n\}$
 iid interarrival times

$$X_n = t_n - t_{n-1}$$

$$0 < E(X) < \infty$$

$$\lambda \stackrel{\text{def}}{=} \frac{1}{E(X)}$$



$$N(t) = \max\{n : t_n \leq t\}$$

$$A(t) = t_{N(t)+1} - t, \quad B(t) = t - t_{N(t)}$$

arrival rate
 $= \lim_{t \rightarrow \infty} \frac{N(t)}{t}$ w.p.1
 ERT

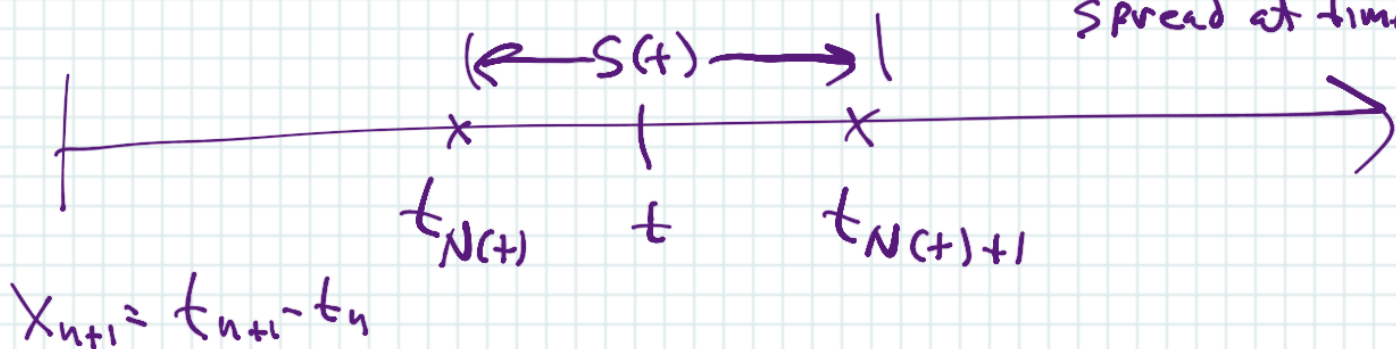
$A(t)$ is called the forward
(excess) recurrence time
at time t

$B(t)$ backwards recurrence time
at time t

(asc)

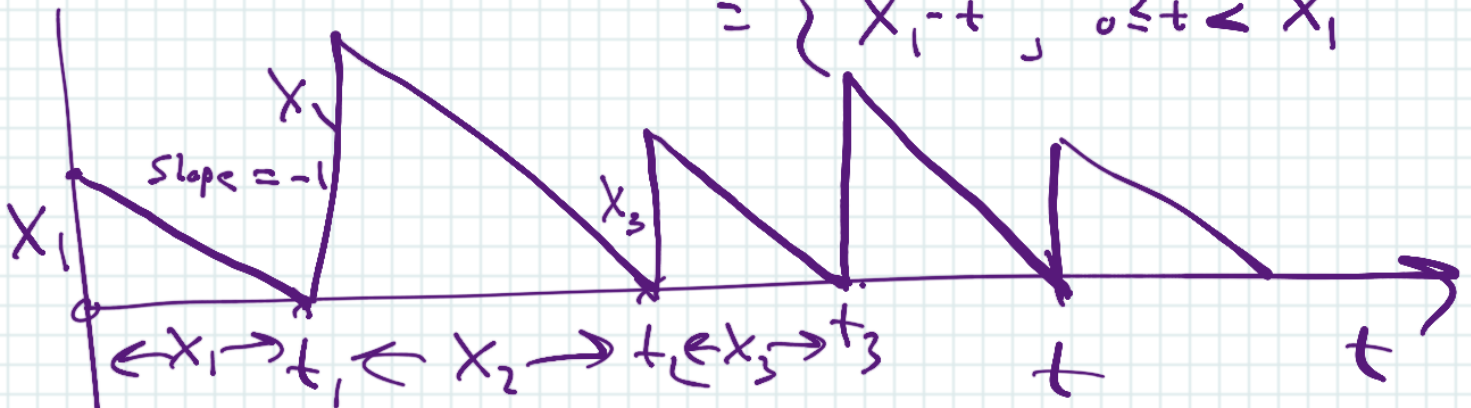
$$S(t) = B(t) + A(t)$$

Spread at time t



$$A(t) = t_{N(t)+1} - t$$

$$= \begin{cases} X_1 - t, & 0 \leq t < X_1 \end{cases}$$



$$\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t A(s) ds$$

$$\int_0^t A(s) ds \approx \frac{X_1^2}{2} + \frac{X_2^2}{2} + \frac{X_3^2}{2} + \frac{X_4^2}{2}$$

$$\left(\frac{1}{t} \sum_{j=1}^{N(t)} \frac{X_j^2}{2} \right) \leq \int_0^t A(s) ds \leq \left(\frac{1}{t} \sum_{j=1}^{N(t)+1} \frac{X_j^2}{2} \right)$$

$$\approx \left(\frac{N(t)}{t} \right) \left(\frac{1}{N(t)} \sum_{j=1}^{N(t)} \frac{X_j^2}{2} \right)$$

ERT \searrow

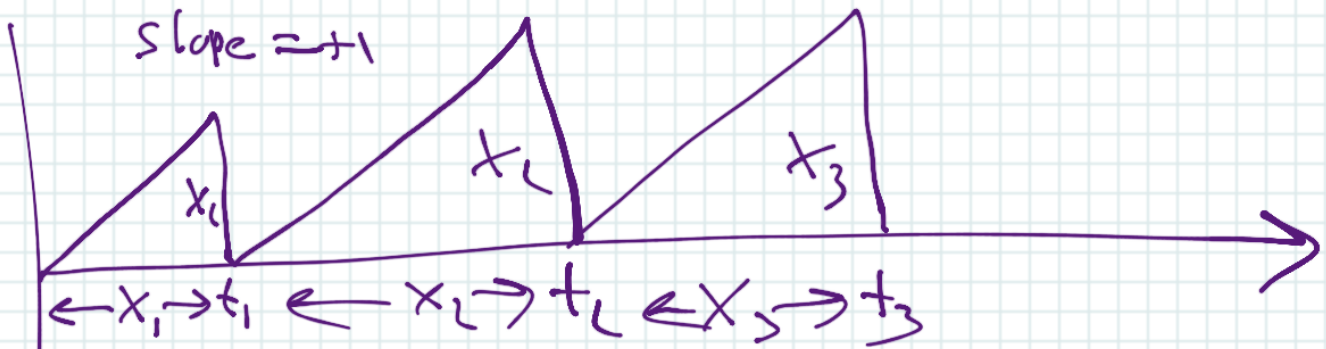
$$\frac{1}{E(X)}$$

\searrow SLLN

$$\frac{E(X^2)}{2}$$

$$\boxed{\frac{E(X^2)}{2 E(X)}}$$

$$B(t) = t - t_{N(t)}$$



$$\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t B(s) ds = \frac{E(X^2)}{2E(X)} \quad \text{wp1}$$

also

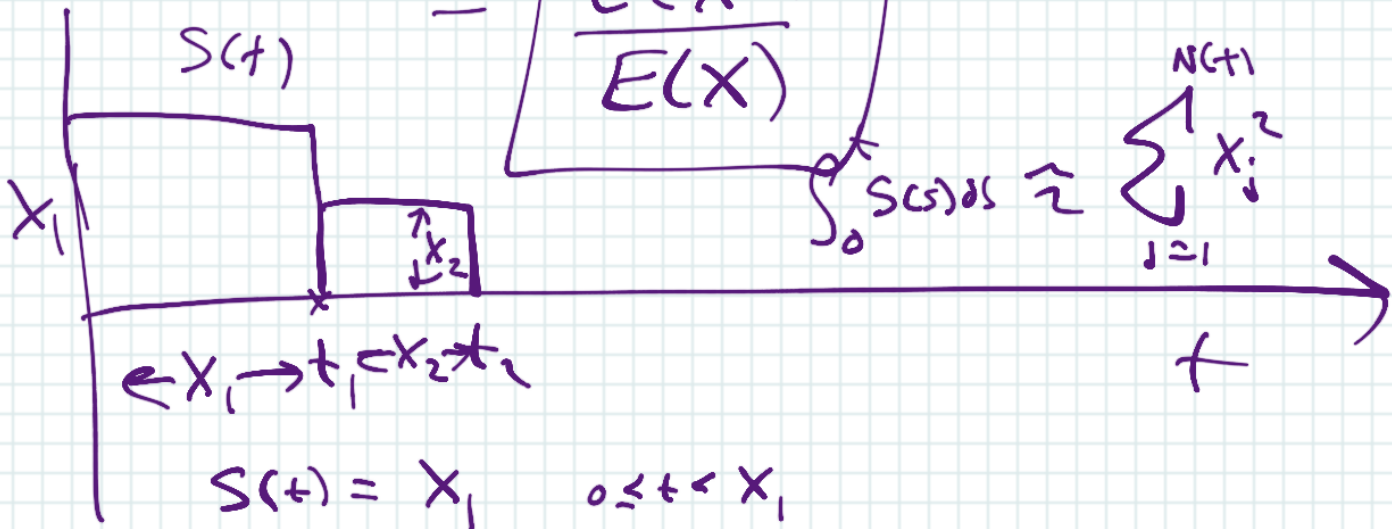
$$\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t S(s) ds$$

$$S(t) = t_{N(t)+1} - t_{N(t)}$$

$$= B(t) + A(t)$$

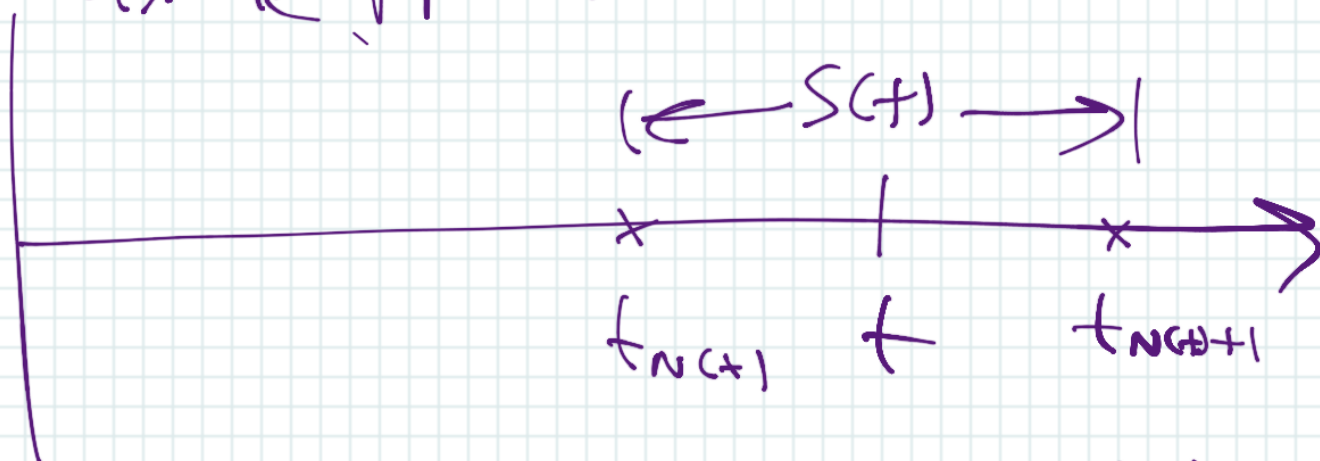
spread at time t

$$= \frac{E(X^2)}{E(X)}$$



assume $PP(\lambda)$

$$S(t) = B(t) + A(t)$$



$$A(t) \sim \exp(\lambda)$$

$$\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t S(s) ds = \frac{E(X^2)}{E(X)} = \frac{2}{\lambda} = 2E(X)$$

$$\lambda = \frac{1}{E(X)}, \quad E(X^2) = \frac{2}{\lambda^2} \quad (X \sim \exp(\lambda))$$

$$\text{Var}(X) = E(X^2) - E^2(X) > 0$$

$$\Rightarrow E(X^2) \geq E^2(X)$$

$$\Rightarrow \frac{E(X^2)}{E(X)} \geq \frac{E^2(X)}{E(X)} = E(X)$$

Inspection paradox

Average interval
you land in is
larger than the
average of (X_i) , $E(X)$

$$\frac{E(X^2)}{E(X)} \geq E(X)$$

$$\{X_n\} = \{1, 2, 1, 1, 1, \dots\}$$

$$\frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{n \rightarrow \infty} \frac{1}{2}(1) + \frac{1}{2}(2) = \frac{3}{2}$$

in time t , $\frac{2}{3}$ of the time q length 2 is in progress

$\frac{1}{3}$ of the time q length 1 is in progress

$$\frac{2}{3}(2) + \frac{1}{3}(1) = \frac{5}{3} > \frac{3}{2}$$

reverse length of the interarrival time you land in $= 5/3$

$$P(X=0) = .99$$

$$P(X=10) = .01$$

$$S(t) = 10$$

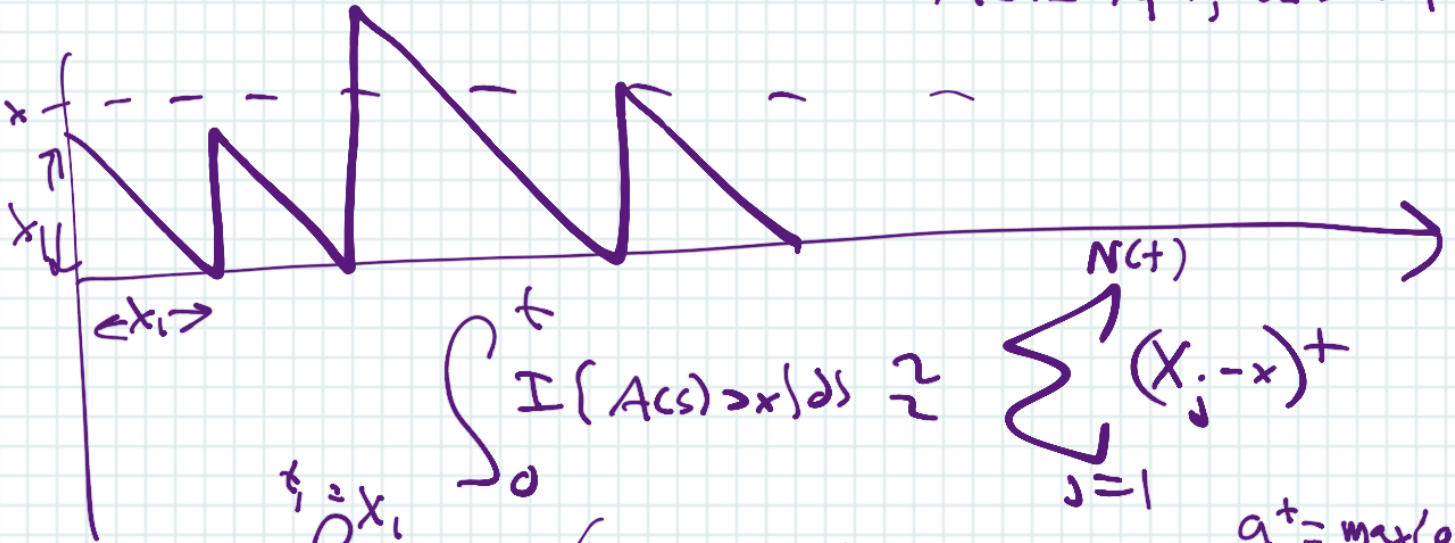
$$E(X) = .99(0) + (.01)(10) = 0.1$$

$$\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t S(s) ds = 10 = \frac{E(X^2)}{E(X)} = \frac{1}{.1} = 10$$

$$E(X^2) = .99(0^2) + .01(10^2) = (.01)(100) = 1$$

$$\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t \mathbf{I}\{A(s) > x\} ds, \quad x \geq 0$$

$$A(s) = X_1 - s, \quad 0 \leq s < X_1$$



$$\int_0^t \mathbf{I}\{A(s) > x\} ds \approx \sum_{j=1}^{N(t)} (X_j - x)^+$$

$$\int_0^{x_1} \mathbf{I}\{x_1 - s > x\} ds = \begin{cases} x_1 - x & \text{if } x_1 > x \\ 0 & \text{if } x_1 \leq x \end{cases}$$

$a^+ = \max\{0, a\}$
Pos. part
of a

$$\frac{1}{t} \int_0^t I(A(s) > x) ds$$

$$\approx \frac{1}{t} \sum_{j=1}^{N(t)} (X_j - x)^+$$

$$\left(\frac{N(t)}{t} \right) \left(\frac{1}{N(t)} \sum_{j=1}^{N(t)} (X_j - x)^+ \right)$$

$t \rightarrow \infty$

ERT $\rightarrow \lambda = \frac{1}{E(t)}$

SLLN $\rightarrow E(X - x)^+$

$$\overline{F}_e(x) \stackrel{\text{def}}{=} \frac{E((X-x)^+)}{E(X)}$$

$$F(x) = P(X \leq x) \quad x \geq 0$$

$$\overline{F}(x) = 1 - F(x) \\ = P(X > x)$$

$$\text{Let } Y = (X-x)^+ \geq 0$$

$$E(Y) = \int_0^{\infty} P(Y \geq y) dy = \int_0^{\infty} P((X-x)^+ \geq y) dy$$

$$= \int_0^{\infty} P(X > y+x) dy = \int_x^{\infty} P(X > y) dy$$

$$\overline{F_e(x)} = \frac{\int_x^\infty \overline{F(y)} dy}{\int_0^\infty \overline{F(y)} dy} \xrightarrow{x \rightarrow \infty} 0$$

$$F_e(x) = \frac{\int_0^x \overline{F(y)} dy}{\int_0^\infty \overline{F(y)} dy} \xrightarrow{x \rightarrow \infty} \frac{E(x)}{E(x)}$$

Defines a proba distri called equilibrium distri

$$\lambda = \frac{1}{E(X)} \quad (0 < E(X) < \infty)$$

$$F_e(x) = \lambda \int_0^x \bar{F}(y) dy, \quad x \geq 0$$

always continuous

$$\int_0^{\infty} n x^{n-1} \bar{F}(x) dx = E(X^n)$$

$$f_e(x) = F_e'(x) = \lambda \bar{F}(x)$$

Let X_e denote a rv dist. as F_e

$$E(X_e) = \int_0^{\infty} x f_e(x) dx = \frac{\lambda}{2} \int_0^{\infty} 2x \bar{F}(x) dx = \frac{\lambda}{2} E(X^2) = \frac{E(X^2)}{2E(X)} \quad n=2$$

example: $X \sim \exp(x)$

$$F(x) = e^{-\lambda x}$$

$$f_e(x) = \lambda e^{-\lambda x}$$

$$\sim \exp(\lambda)$$

example $P(X=1) = 1$ $\lambda = 1$

$$F(x) = \begin{cases} 1, & 0 \leq x < 1 \\ 0, & x \geq 1 \end{cases}$$

$$f_e(x) = \begin{cases} 1 & x \in (0, 1) \\ 0 & x \notin (0, 1) \end{cases}$$

unif(0,1) density