

Applied Mathematics Reading Seminar

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- Introduction
- Ensemble Adjustment Kalman Filter
- Adaptive Inflation
- Numerical Results

- Time-varying reproduction number R_t is the average number of secondary cases caused by an infectious individual in a well-mixed population of susceptible individuals at time t
- Online estimation of time-varying parameters (e.g., R_t) can identify periods of heightened transmission and inform control strategies
- Current data assimilation methods often face challenges with **underestimating uncertainty** and capturing temporal trends when there are abrupt changes in parameters

Covariance Inflation

State Space Model

Dynamics models: $v_{n+1} = \Psi(v_n) + \xi_n, n \in \{1, \dots, N\}$

Data Model: $y_{n+1} = h(v_{n+1}) + \eta_{n+1}, n \in \{1, \dots, N\}$

Probabilistic Structure: $v_0 \sim \mathcal{N}(m_0, C_0), \xi_n \sim \mathcal{N}(0, \Sigma), \eta_n \sim \mathcal{N}(0, \Gamma)$ i.i.d.

Probabilistic Structure: $v_0 \perp \{\xi_n\} \perp \{\eta_n\}$

States (and parameters): $v_n \in \mathbb{R}^d$. *Observations:* $y_n \in \mathbb{R}^k$.

Assume $C_0, \Sigma, \Gamma > 0, \Psi \in C(\mathbb{R}^d, \mathbb{R}^d)$, and $h \in C(\mathbb{R}^d, \mathbb{R}^k)$ are all known.

Observed data:

$$Y_n^\dagger := \{y_1^\dagger, \dots, y_n^\dagger\}$$

Goal: iteratively compute posterior distribution

$$\mathbb{P}(v_n | Y_n^\dagger), n = 1, 2, \dots, N$$

Represent distribution with an ensemble of particles $\{v_n^{(j)}\}_{j=1}^J$ so

$$\mathbb{P}(v_n | Y_n^\dagger) \approx \frac{1}{J} \sum_{j=1}^J \delta(v_n - v_n^{(j)}).$$

Bayes Rule

$$\mathbb{P}(v_n | Y_n^\dagger) \propto \underbrace{\mathbb{P}(v_n | Y_{n-1}^\dagger)}_{\text{prior}} \underbrace{\mathbb{P}(y_n | v_n)}_{\text{likelihood}}$$

(1)

$$\text{Forecast} \left\{ \begin{array}{l} \hat{v}_{n+1}^{(j)} = \Psi \left(v_n^{(j)} \right) + \xi_n^{(j)}, j = 1, \dots, J, \\ \hat{m}_{n+1} = \frac{1}{J} \sum_{j=1}^J \hat{v}_{n+1}^{(j)}, \quad \hat{C}_{n+1} = \frac{1}{J-1} \sum_{j=1}^J \left(\hat{v}_{n+1}^{(j)} - \hat{m}_{n+1} \right) \left(\hat{v}_{n+1}^{(j)} - \hat{m}_{n+1} \right)^\top \\ \hat{H}_{n+1} = \frac{1}{J-1} \sum_{j=1}^J \left(h(\hat{v}_{n+1}^{(j)}) - \sum_{j=1}^J h(\hat{v}_{n+1}^{(j)}) \right) \left(h(\hat{v}_{n+1}^{(j)}) - \sum_{j=1}^J h(\hat{v}_{n+1}^{(j)}) \right)^\top \end{array} \right.$$

(2)

$$\text{Analysis} \left\{ \begin{array}{l} K_{n+1} = \hat{C}_{n+1}^{\frac{1}{2}} \hat{H}_{n+1}^{\frac{1}{2}\top} \left(\hat{H}_{n+1} + \Gamma \right)^{-1}, \\ m_{n+1} = \hat{m}_{n+1} + K_{n+1} \left(y_{n+1}^\dagger - \frac{1}{J} \sum_{j=1}^J h(\hat{v}_{n+1}^{(j)}) \right), \\ C_{n+1} = \hat{C}_{n+1}^{\frac{1}{2}} \left(I + \hat{H}_{n+1}^\top \Gamma^{-1} \hat{H}_{n+1} \right)^{-1} \hat{C}_{n+1}^{\frac{1}{2}\top} \\ v_{n+1}^{(j)} = m_{n+1} + C_{n+1}^{\frac{1}{2}} \hat{C}_{n+1}^{-\frac{1}{2}} \left(\hat{v}_{n+1}^{(j)} - \hat{m}_{n+1} \right) \end{array} \right.$$

Covariance Inflation

To correct the underestimated the uncertainty of ensemble $\{v_n^{(j)}\}_{j=1}^J$ we can inflate the variance by a factor λ_n

$$v_n^{(j)} \leftarrow \sqrt{\lambda_n}(v_n^{(j)} - m_n) + m_n$$

- No inflation $\lambda_n = 1$
- Fixed inflation $\lambda_n = 1.02$
- Adaptive inflation $\lambda_n = \underset{\lambda}{\operatorname{argmax}} p(\lambda | Y_n^\dagger)$

Covariance Inflation

Laplace Approximation

$$\begin{aligned} p(\lambda | y_n) &\sim N(\mu, \Sigma) \\ p(\lambda | y_n) &= \frac{p(y_n, \lambda)}{\int p(y_n, \lambda) d\lambda} \\ &= \frac{e^{\ln p(y_n, \lambda)}}{\int e^{\ln p(y_n, \lambda)} d\lambda} \end{aligned}$$

Choose $\lambda_0 = \lambda_{MAP}$.

$$\begin{aligned} p(y_n, \lambda) &\approx p(y_n, \lambda_{MAP}) + (\lambda - \lambda_0)^\top \nabla p(y_n, \lambda_{MAP}) + \\ &\quad \frac{1}{2} (\lambda - \lambda_0)^\top \nabla^2 p(y_n, \lambda_{MAP}) (\lambda - \lambda_0) \end{aligned}$$

Adaptive Inflation

See board...

S(Susceptible) I(Infected) R(Recovered) N (fixed population)

$$(3) \quad \begin{aligned} \frac{dS}{dt} &= -\frac{\beta(t)SI}{N} \\ \frac{dI}{dt} &= \frac{\beta(t)SI}{N} - \gamma I \\ \frac{dR}{dt} &= \gamma I \end{aligned}$$

$$\beta(t) = \beta_0 + \frac{\beta_1 - \beta_0}{1 + e^{-k(t-m)}}$$

Stochastic SIR Model

S(Susceptible) I(Infected) R(Recovered)

$$dSI \sim \text{Pois} \left(\frac{\beta_t S_t I_t}{N} \right)$$

$$dIR \sim \text{Pois}(\gamma I_t)$$

$$(4) \quad \left. \begin{aligned} S_{t+1} &= S_t - dSI \\ I_{t+1} &= I_t + dSI - dIR \\ R_{t+1} &= R_t + dIR \\ i_{t+1} &= dSI \end{aligned} \right\} \psi$$

$$v_t = \{S_t, I_t, \gamma, \beta_t, i_t\}. \quad h = H = [0, 0, 0, 0, 1]^\top.$$

Numerical Results

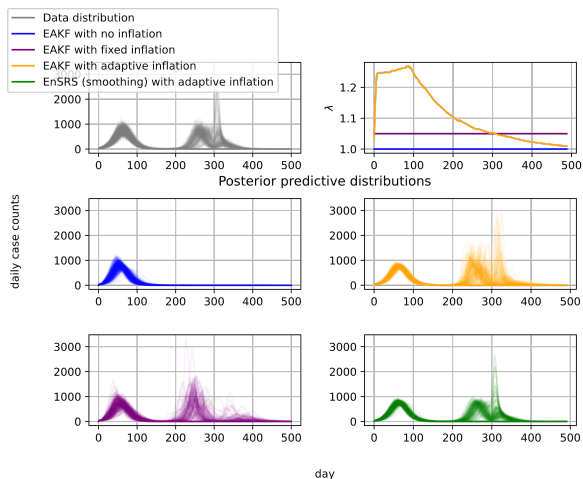


Figure 1: Underlying data distribution, inflation λ , and posterior predictive distributions for each of the 4 methods. Each line in the posterior predictive distribution represents one realization.

Numerical Results

Ran each method for 100,000 parameter scenarios with 100 realizations for each method.

Method	$D_{KL}(p q)$	$W_2(p, q)$	Prop of true R_t within 95% CI
EAKF no inflation	7.89	1.08	0.00
EAKF fixed inflation	1.75	0.72	0.15
EAKF adaptive inflation	0.63	0.52	0.97
EnSRS adaptive inflation	0.50	0.39	0.97

Table 1: Performance metrics for posterior predictive distributions p and data distribution q . Proportion of true R_t within 95% CI was computed on the day the sigmoid curve had reached 99% of its final value after the last change point.

- PAPER
- Investigate other data assimilation methods
- Adaptive assimilation window