

**Midterm Solution.** Fall 2005.  
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1. The state space of the Markov chain is  $\{(i, j) | i \geq 1, j \geq 1\}$ . Suppose that the Markov chain is currently in the state  $(i, j)$ . Let  $T$  denote the time the Markov chain stays in state  $(i, j)$ , and let  $T_{kl}$  denote the time until the  $k$ -th male mates with the  $l$ -th female. Since they mate in any time interval of length  $h$  with probability  $\lambda h + o(h)$ , it follows that  $T_{kl}$  is exponentially distributed with rate  $\lambda$ . Since

$$T = \min_{1 \leq k \leq i, 1 \leq l \leq j} T_{kl}$$

and the  $ij$  random variables  $T_{kl}$  are independent with each other, it follows that  $T$  is exponentially distributed with rate  $ij\lambda$ , namely

$$v_{(i,j)} = ij\lambda.$$

Since each offspring is equally likely to be male or female, we conclude

$$P_{(i,j)(i,j+1)} = P_{(i,j)(i+1,j)} = \frac{1}{2}.$$

2. (a). Three. They are  $\{0,1\}$ ,  $\{2\}$ , and  $\{3\}$ .  $\{2\}$  is a transient state. The others are recurrent.

(b).

$$\pi_0 = .7\pi_0 + .25\pi_1$$

$$\pi_1 = .3\pi_0 + .75\pi_1$$

$$\pi_0 + \pi_1 = 1.$$

This yields  $\pi_0 = \frac{5}{11}$ ,  $\pi_1 = \frac{6}{11}$ .

(c).  $p = 0.10 * 1 + 0.20 * 1 + 0.40 * p + 0.30 * 0$ . Thus,  $p = 1/2$ .

(d).

$$\lim_{n \rightarrow \infty} p_{20}^{(n)} = p\pi_0 = \frac{1}{2} * \frac{5}{11} = \frac{5}{22}$$

$$\lim_{n \rightarrow \infty} p_{21}^{(n)} = p\pi_1 = \frac{1}{2} * \frac{6}{11} = \frac{6}{22}.$$

(e).

$$\lim_{n \rightarrow \infty} p_{00}^{(n)} = \pi_0 = \frac{5}{11}, \quad \lim_{n \rightarrow \infty} p_{10}^{(n)} = \pi_0 = \frac{5}{11},$$

$$\lim_{n \rightarrow \infty} p_{01}^{(n)} = \pi_1 = \frac{6}{11}, \quad \lim_{n \rightarrow \infty} p_{11}^{(n)} = \pi_1 = \frac{6}{11}.$$

Since 0 and 1 can only make transitions between 0 and 1, we have

$$\lim_{n \rightarrow \infty} p_{02}^{(n)} = \lim_{n \rightarrow \infty} p_{03}^{(n)} = \lim_{n \rightarrow \infty} p_{12}^{(n)} = \lim_{n \rightarrow \infty} p_{13}^{(n)} = 0$$

Since 2 is a transient state, we have

$$\lim_{n \rightarrow \infty} p_{22}^{(n)} = 0,$$

and

$$\lim_{n \rightarrow \infty} p_{23}^{(n)} = 1 - \lim_{n \rightarrow \infty} p_{20}^{(n)} - \lim_{n \rightarrow \infty} p_{21}^{(n)} - \lim_{n \rightarrow \infty} p_{22}^{(n)} = 1 - \frac{5}{22} - \frac{6}{22} - 0 = \frac{1}{2}.$$

$\lim_{n \rightarrow \infty} p_{20}^{(n)}$  and  $\lim_{n \rightarrow \infty} p_{21}^{(n)}$  are given in Part (d). Since 3 is an absorbing state, we have

$$\lim_{n \rightarrow \infty} p_{03}^{(n)} = \lim_{n \rightarrow \infty} p_{13}^{(n)} = \lim_{n \rightarrow \infty} p_{23}^{(n)} = 0, \quad \lim_{n \rightarrow \infty} p_{33}^{(n)} = 1.$$

In summary, we have

$$\lim_{n \rightarrow \infty} \mathbf{P}^{(n)} = \begin{pmatrix} 5/11 & 6/11 & 0 & 0 \\ 5/11 & 6/11 & 0 & 0 \\ 5/22 & 6/22 & 0 & 1/2 \\ 0 & 0 & 0 & 1 \end{pmatrix}.$$

3. Let  $A$  be the event that machine 1 is still working at time  $t$ . Then

$$P(A) = e^{-\lambda_1 t}, \quad P(A^c) = 1 - e^{-\lambda_1 t}.$$

Furthermore,

$$P(\text{machine 1 is the first machine to fail} | A^c) = 1.$$

In addition by the memoryless property,

$$\begin{aligned} & P(\text{machine 1 is the first machine to fail} | A) \\ &= P(\text{machine 1 fails before machine 2}) \\ &= \frac{\lambda_1}{\lambda_1 + \lambda_2}. \end{aligned}$$

In summary, we have

$$\begin{aligned} & P(\text{machine 1 is the first machine to fail}) \\ = & P(\text{machine 1 is the first machine to fail}|A)P(A) \\ & + P(\text{machine 1 is the first machine to fail}|A^c)P(A^c) \\ = & \frac{\lambda_1}{\lambda_1 + \lambda_2}e^{-\lambda_1 t} + 1 - e^{-\lambda_1 t}. \end{aligned}$$