

USING SIMULATION TO STUDY SERVICE-RATE CONTROLS TO STABILIZE PERFORMANCE IN A SINGLE-SERVER QUEUE WITH TIME-VARYING ARRIVAL RATE

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December 5, 2015

Outline

- 1 Motivation
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 - Service-Rate Controls
- 2 The Model
- 3 Simulation Methods For Nonstationary Models
 - Generating the Arrival Process
 - Generating the Service Times
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Motivation

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M. Defraeye and I. Van Nieuwenhuysse (2013) Controlling excessive waiting times in small service systems with time-varying demand: an extension of the ISA algorithm. *Decision Support Systems* 54(4), 1558 – 1567.



Y. Liu and W. Whitt (2012) Stabilizing customer abandonment in many-server queues with time-varying arrivals. *Oper. Res.* 60(6), 1551 – 1564.



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Example (use a service-rate control)

- 1 Hospital Surgery Rooms
- 2 Airport Security Inspection Lines

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We use simulation to study service-rate controls to stabilize performance in a single-server queue with time-varying arrival rates.

- We conduct simulation experiments to **evaluate the performance** of alternative service-rate controls.
- We **develop an efficient algorithm** for simulating a time-varying queue with a service-rate control.

The Model

The $G_t/G_t/1$ queue

$G_t/G_t/1$ Single-Server Queueing Model

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- Unlimited waiting space
- Service rate is subject to control
- i.i.d. service requirements separate from the service rate

The Arrival Process

The Arrival Process

A **time-transformation** of a stationary counting process:

$$A(t) \equiv N_a(\Lambda(t)) \equiv N_a\left(\int_0^t \lambda(s) ds\right), \quad t \geq 0, \quad (1)$$

where

- Λ is the **cumulative** arrival rate function:

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- All the stochastic variability is separated from the deterministic arrival rate function.

The Service Process

Queue Length and Departure Process

$$Q(t) \equiv A(t) - D(t), \quad t \geq 0, \quad (2)$$

$$D(t) \equiv N_s \left(\int_0^t \mu(s) 1_{\{Q(s) > 0\}} ds \right), \quad t \geq 0, \quad (3)$$

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- The service requirement process N_s is separated from the deterministic service-rate function $\mu(t)$.

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$$E[W(t)] \approx \rho(t)V / \mu(t)(1 - \rho(t)) = \lambda(t)V / (\mu(t)^2 - \mu(t)\lambda(t)).$$

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- **Problem:** need to compute Λ^{-1} for each arrival in each simulation run.
- Compute the inverse function Λ^{-1} for one cycle outside of simulation and do table lookup when simulating.
 - $\Lambda^{-1}(kC + t) = kC + \Lambda^{-1}(t)$ for $0 \leq t \leq C$, where C is the length of a cycle.
- (See the paper for the details.)

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- **But V_k is not formulated.**

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Exact service time formula:

$$S_k = \int_{B_k}^{B_k + V_k} \mu(s) ds, \quad k \geq 1. \quad (7)$$

If we let

$$M(t) \equiv \int_0^t \mu(s) ds, \quad t \geq 0, \quad (8)$$

then

$$V_k = M^{-1}(S_k + M(B_k)) - B_k, \quad k \geq 1. \quad (9)$$

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Simulation Experiments

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- $\beta=0.2$, $\gamma=0.001, 0.01, 0.1, 1, 10$.
- To cover a range of difference cycle lengths of $2\pi/\gamma$.

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- Erlang (sum of two i.i.d. exponentials, $c^2 = 0.5$)

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Simulation Results

Simulation Results: The Rate-Matching Control

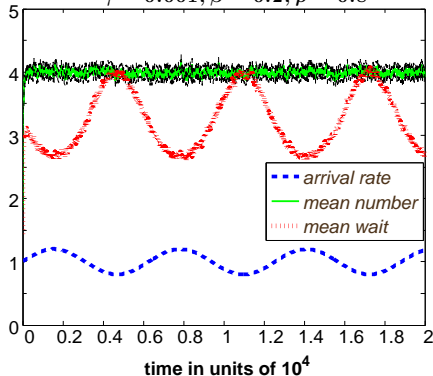
1. $\gamma=0.001$

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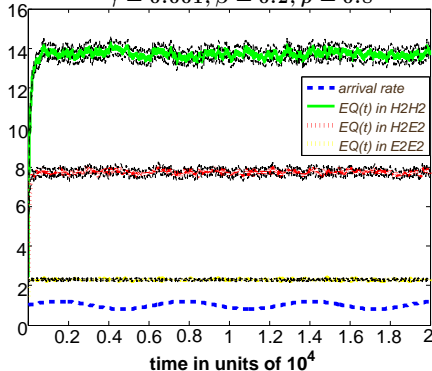
1. $\gamma=0.001$

- Cycle length is 6.28×10^3 .
- The left graph is **Markovian model**; the right graph shows (H_2/H_2) , (H_2/E_2) and (E_2/E_2) .
- $E(Q(t))$ stabilized at target, but $E(W(t))$ is periodic.

$\gamma = 0.001, \beta = 0.2, \rho = 0.8$



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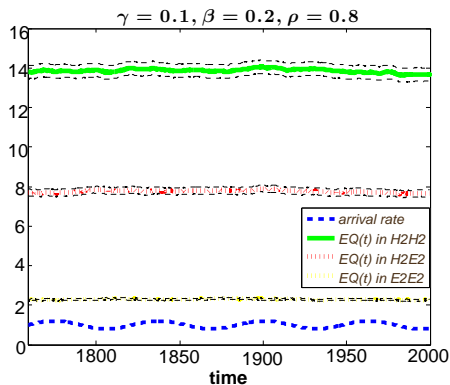
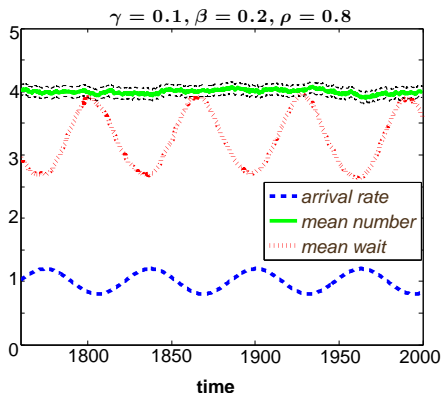
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- Cycle length is 62.8, only last 3 to 4 cycles are displayed.
- $E(Q(t))$ stabilized at target, but $E(W(t))$ is periodic.



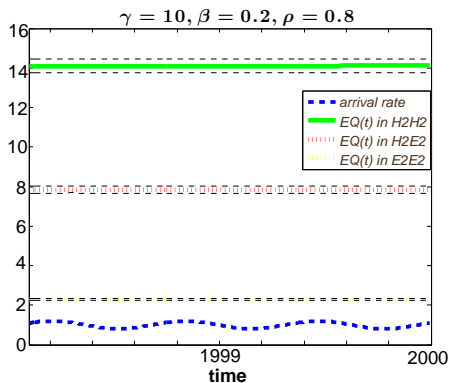
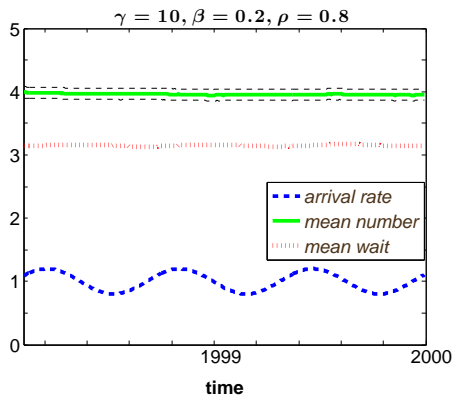
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3. $\gamma=10$

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- Cycle length is 0.63, only last 3 cycles are displayed.
- By Whitt (1984) the system converges to stationary case.



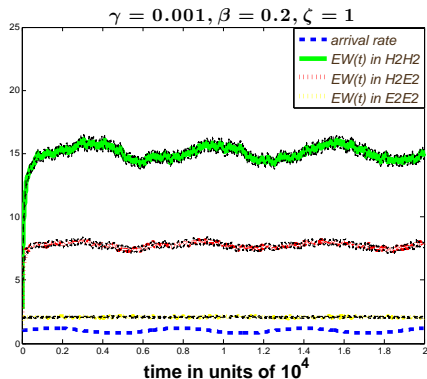
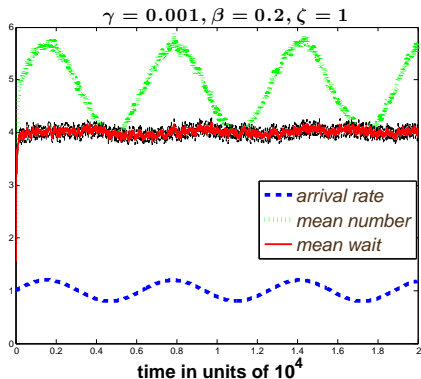
Simulation Results: The Square-Root Control

1. $\gamma=0.001$

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- Cycles are long, and arrival rates change slowly, thus PSA is appropriate. [Whitt, 1991]
- $E(W(t))$ is stabilized, while $E(Q(t))$ is periodic.



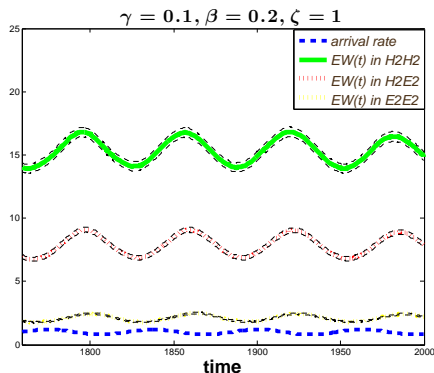
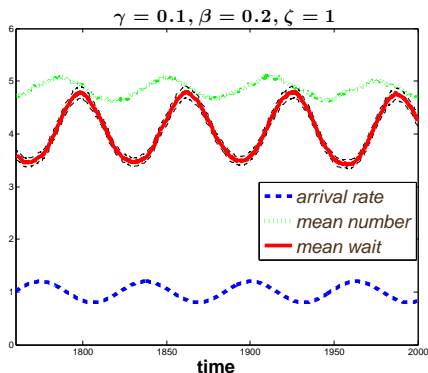
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2. $\gamma=0.1$

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- PSA does not hold as cycles are short.
- Neither $E(W(t))$ nor $E(Q(t))$ is stabilized.



Thank You!

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- Calculate Λ value for n_x equally spaced points of $[0, C]$, let spacing be $\eta = C/n_x$.
- Construct approximation J of Λ^{-1} over n_y equally spaced points in $[0, C]$, let spacing be $\delta = C/n_y$.
 - Using $J(j\delta) = k\eta$, $1 \leq j \leq n_y$,
where $0 \leq k \leq n_x$ and $k\eta$ is closest point greater equal to the true inverse value.

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 - Using $J(j\delta) = k\eta$, $1 \leq j \leq n_y$,
where $0 \leq k \leq n_x$ and $k\eta$ is closest point greater equal to the true inverse value.
- Extend J to $[0, C]$ by letting $J(t) = J(\lfloor t/\delta \rfloor \delta)$.

Results From [Whitt, 2015]: The Rate Matching Control

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Theorem 2.1 (*time transformation of a stationary model*)

For (A, D, Q) with the rate-matching service-rate control and the stationary single-server model (A_1, D_1, Q_1) ,

$$(A(t), D(t), Q(t)) = (A_1(\Lambda(t)), D_1(\Lambda(t)), Q_1(\Lambda(t))), \quad t \geq 0. \quad (11)$$

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Theorem 3.2 (*stabilizing the queue-length distribution and the steady-state delay probability*)

Let $Q_1(t)$ be the queue length process when $\lambda(t) = 1$, $t \geq 0$. If $Q_1(t) \Rightarrow Q_1(\infty)$ as $t \rightarrow \infty$, where $P(Q_1(\infty) < \infty) = 1$, then also

$$Q(t) \Rightarrow Q_1(\infty) \quad \text{in } \mathbb{R} \quad \text{as } t \rightarrow \infty, \quad (12)$$

and

$$P(W(t) > 0) = P(Q(t) \geq 1) \rightarrow \rho \quad \text{as } t \rightarrow \infty. \quad (13)$$

Results From [Whitt, 2015]: The Square-Root Control

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Section 6.3 (*stabilizing the expected time-varying virtual waiting time*)

We assume that the Pointwise Stationary Approximation (PSA) is appropriate. Then the square-root control (14) stabilizes $E[W(t)]$ at the target w for all t under heavy traffic.

$$\mu(t) \equiv \lambda(t) + \frac{\lambda(t)}{2} \left(\sqrt{1 + \frac{\zeta}{\lambda(t)}} - 1 \right), \quad t \geq 0, \quad (14)$$

where ζ is inversely proportional to w .

Pointwise Stationary Approximation (PSA)

Performance at different times can be regarded as approximately the same as the performance of the stationary system with the model parameters operating at those separate times.

Theorem From [Whitt, 2015]

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Theorem 6.1 (impossibility of stabilizing both the waiting time distribution and the mean number in queue)

Consider a $G_t/G_t/1$ system starting empty in the distant past. Suppose that a service-rate control makes $P(W(t) > x)$ independent of t for all $x \geq 0$. Then the only arrival rate functions for which the mean number waiting in queue $E[(Q(t) - 1)^+]$ is constant, independent of t , are the constant arrival rate functions.

Theorem From [Whitt, 2015]

Theorem 6.1 (impossibility of stabilizing both the waiting time distribution and the mean number in queue)

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Theorem 1 (*Convergence of point processes*)

The point process $D(t)$ has predictable stochastic intensity $\Lambda(t)$, then it can be represented as the random-time transformation

$$D(t) = \Pi(C(t)), \quad t \geq 0, \quad (15)$$

where $\Pi(t)$ is a Poisson process with unit intensity and $C(t) = \int_0^t \Lambda(u) du$. If $C_n(t) \Rightarrow ct$ in \mathbb{R} as $n \rightarrow \infty$ for each t , then $D_n \Rightarrow \Pi_c$ in $D[0, \infty)$ as $n \rightarrow \infty$, where Π_c is a Poisson process with intensity c .

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Section 3.2 (*Applications to queues*)

Apply rescaling

$$D_n(t) = \hat{D}_n(t)(t/n) \quad \text{and} \quad C_n(t) = \hat{C}_n(t)(t/n) \quad (16)$$

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